Augenblick, Lazarus: Restrictions on Asset-Price Movements Under Rational Expectations: Theory and Evidence

Discussion by Jaroslav Borovička (NYU) July 2018

- 1. Beliefs, risk adjustments and asset prices
 - · information in no-arbitrage restrictions
- 2. Summary of the paper methodology and results
 - martingale restrictions on beliefs
 - $\cdot\,$ localization of state prices
 - implied slopes of the SDF
- 3. Role of assumptions
 - key assumption: constant SDF slope
- 4. Questions and suggestions
 - are we testing for rational expectations?

High volatility in prices of risky securities

- high frequency (intraday, daily returns)
- business cycle frequency (monthly, quarterly returns)

Are these asset price fluctuations 'rational'?

• i.e., can they be explained by a plausible model of investors' preferences, market constraints, and investors' beliefs that are correct (in line with the data-generating process (DGP))?

BACKGROUND

Abstract from market constraints, assume no arbitrage.

• then, for any investor's subjective belief \hat{P} , there exists an SDF m_{t+1} that prices financial assets

$$1 = \widehat{E}_t \left[m_{t+1} R_{t+1} \right]. \tag{1}$$

BACKGROUND

Abstract from market constraints, assume no arbitrage.

• then, for any investor's subjective belief \hat{P} , there exists an SDF m_{t+1} that prices financial assets

$$1 = \widehat{E}_t \left[m_{t+1} R_{t+1} \right]. \tag{1}$$

• but of course, defining $h = d\hat{P}/dP$, we can write

$$1 = E_t[\underbrace{m_{t+1}h_{t+1}}_{\widehat{m}_{t+1}}R_{t+1}]$$
(2)

BACKGROUND

Abstract from market constraints, assume no arbitrage.

• then, for any investor's subjective belief \widehat{P} , there exists an SDF m_{t+1} that prices financial assets

$$1 = \widehat{E}_t \left[m_{t+1} R_{t+1} \right]. \tag{1}$$

• but of course, defining $h = d\widehat{P}/dP$, we can write

$$1 = E_t[\underbrace{m_{t+1}h_{t+1}}_{\widehat{m}_{t+1}}R_{t+1}]$$
(2)

Which is the right model of the SDF, m_{t+1} or \hat{m}_{t+1} ?

- \cdot that depends on what is plausible \implies impose structure on the SDF
- asset price data alone cannot tell (1) and (2) apart
 - even if we perfectly know the DGP measure P

Information from no-arbitrage restrictions

$$1 = E_t[\underbrace{m_{t+1}h_{t+1}}_{\widehat{m}_{t+1}}R_{t+1}]$$

- Time series information identifies the DGP measure P.
- Data on returns R_{t+1} together with no-arbitrage restrictions inform \widehat{m}_{t+1} .
- We must impose additional restrictions to separately infer something about m_{t+1} and h_{t+1} .

This paper

$$1 = E_t[\underbrace{m_{t+1}h_{t+1}}_{\widehat{m}_{t+1}}R_{t+1}]$$

- imposes the zero of correct beliefs $\implies h_{t+1} \equiv 1$
- restricts the class of SDFs (conditional transition independence (CTI)) \implies restriction on m_{t+1}
- uses martingale restrictions under DGP (time series information) to deduce whether $\exists m_{t+1}$ in the given class that also prices returns

This paper

$$1 = E_t[\underbrace{m_{t+1}h_{t+1}}_{\widehat{m}_{t+1}}R_{t+1}]$$

- imposes the zero of correct beliefs $\implies h_{t+1} \equiv 1$
- restricts the class of SDFs (conditional transition independence (CTI)) \implies restriction on m_{t+1}
- uses martingale restrictions under DGP (time series information) to deduce whether $\exists m_{t+1}$ in the given class that also prices returns

Outcomes

- if \exists such m_{t+1} , characterize the minimum needed local risk aversion consistent with $h_{t=1} \equiv 1$
- if \nexists such m_{t+1} , conclude that $h_{t+1} \neq 1$.

MARTINGALE RESTRICTIONS

Let X_t be a square-integrable martingale. Then

$$E_t\left[\sum_{j=t}^{T-1} (X_{j+1} - X_j)^2\right] = E_t\left[X_T^2 - X_t^2\right]$$

 \cdot restriction on quadratic variation of a martingale

MARTINGALE RESTRICTIONS

Let X_t be a square-integrable martingale. Then

$$E_t\left[\sum_{j=t}^{T-1} (X_{j+1} - X_j)^2\right] = E_t\left[X_T^2 - X_t^2\right]$$

restriction on quadratic variation of a martingale

'Beliefs are martingales' (under their own measure)

- \cdot apply the above result to a two-state outcome \implies localization
- $\cdot \pi_t$ probability of the low state realization at T

$$E_{t}[m_{T}] \doteq E_{t}\left[\sum_{j=t}^{T-1} (\pi_{j+1} - \pi_{j})^{2}\right] = E_{t}[(1 - \pi_{t})\pi_{t} - (1 - \pi_{T})\pi_{T}] \doteq E_{t}[r_{T}]$$

- if we could measure beliefs, this would constitute a test of rational expectations (Augenblick and Rabin (2018))
- · restriction on how much beliefs can vary over time

Instead, we observe risk-neutral (Arrow–Debreu) prices π_t^* . Then

$$E_t [\mathbf{m}_T^*] \doteq E_t \left[\sum_{j=t}^{T-1} \left(\pi_{j+1}^* - \pi_j^* \right)^2 \right] \neq E_t [(1 - \pi_t^*) \pi_t^* - (1 - \pi_T^*) \pi_T^*] \doteq E_t [\mathbf{r}_T^*]$$

Relationship between π_t^* and π_t

$$\pi_t^* \propto E_t \left[M_T / M_t \mid s_T = \text{low} \right] \pi_t$$

· reflects risk adjustment and normalization by the risk-free rate

Rewrite relation between π_t^* and π_t in terms of odds ratio

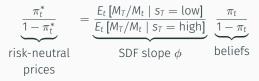
$$\underbrace{\frac{\pi_t^*}{1-\pi_t^*}}_{\text{risk-neutral prices}} = \underbrace{\frac{E_t \left[M_T/M_t \mid s_T = \text{low}\right]}{E_t \left[M_T/M_t \mid s_T = \text{high}\right]}}_{\text{SDF slope } \phi} \underbrace{\frac{\pi_t}{1-\pi_t}}_{\text{beliefs}}$$

If we observe (a lot of) variation in π_t^* over time, what can we infer about variation in π_t ?

- we must restrict movements in the SDF slope ϕ over time
- + ϕ can differ across terminal state realizations

Assumption: Conditional transition independence (CTI)

 $\cdot \phi$ is constant over time (it can also follow a martingale)



Idea: A high ϕ translates large variation in π_t^* into modest variation in π_t , which can then satisfy the martingale restriction on correct beliefs.

$$E_t[m_T^* - r_t^*] \le (\pi_t^*)^2 \left(1 - \frac{1}{\pi_t^* + \phi(1 - \pi_t^*)}\right)$$

Define states: Terminal values of a market index $R_T^m = s_T$ at some date *T*. Extract Arrow–Debreu prices (Breeden, Litzenberger (1978)) Construct sample averages for $E_t [m_T^* - r_t^*]$

- daily, weekly, monthly frequencies
- conditional on terminal returns for various adjacent payoff states $R_T^m \in \{s_j, s_{j+1}\} \implies$ localization of the return distribution

Infer the required ϕ 's to satisfy inequality

$$E_t[m_T^* - r_t^*] \le (\pi_t^*)^2 \left(1 - \frac{1}{\pi_t^* + \phi(1 - \pi_t^*)}\right)$$

Required SDF slopes ϕ need to be very high to rationalize correct beliefs.

• Consequence of high volatility in π_t^* .

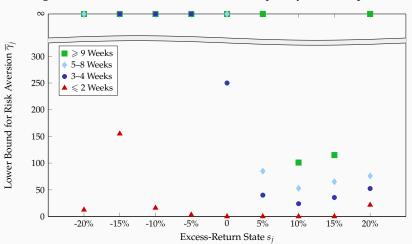


Figure 5: Estimates of Relative Risk Aversion: Splits by Time to Expiration

Required SDF slopes ϕ need to be very high to rationalize correct beliefs.

• Consequence of high volatility in π_t^* .

Higher ϕ needed to rationalize higher frequency (daily) variation.

• Is there more 'unexplained' variation at higher frequencies? Unclear, only a necessary bound.

SUMMARY OF THE RESULTS

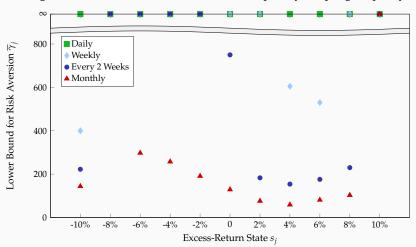


Figure 6: Estimates of Relative Risk Aversion: Splits by Sampling Frequency

Required SDF slopes ϕ need to be very high to rationalize correct beliefs.

• Consequence of high volatility in π_t^* .

Higher ϕ needed to rationalize higher frequency (daily) variation.

• Is there more 'unexplained' variation at higher frequencies? Unclear, only a necessary bound.

High risk aversion needed across all terminal payoff states.

• Informative about the type of risks faced by investors (e.g., tail risk).

SUMMARY OF THE RESULTS

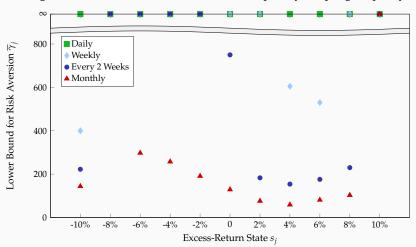


Figure 6: Estimates of Relative Risk Aversion: Splits by Sampling Frequency

DISCUSSION OF ASSUMPTIONS

The whole identification hinges on the choice of class of 'plausible' SDFs.

$$\frac{\pi_t^*}{1-\pi_t^*} = \underbrace{\frac{E_t \left[M_T/M_t \mid R_T^m = \text{low}\right]}{E_t \left[M_T/M_t \mid R_T^m = \text{high}\right]}}_{\text{Linear formula}} \frac{\pi_t}{1-\pi_t}$$

constant SDF slope ϕ $\phi = \frac{u'(R_T^m = \text{low})}{u'(R_T^m = \text{high})}$

To first-order approximation, this corresponds to constant local relative risk aversion $\gamma(\bar{R})$

market risk premium

$$E_t\left[R_T^e\right] = \gamma\left(\overline{R}\right) Var_t\left[R_T^e\right]$$

- $\gamma(\bar{R})$ is the local curvature of the utility/value function as a function of the return realization, by assumption independent of t
- time-variation in market risk premium can only be driven by $Var_t[R_T^e]$

Separable utility over returns at terminal state

• caveat: if utility is over consumption, then time variation in the mapping between returns and consumption can lead to violations

Epstein–Zin in iid growth environment

 \cdot not very interesting, isomorphic to recalibrated separable CRRA utility

Disaster risk model (Gabaix (2012)) with CRRA utility

- holds for (high) return realizations conditional on which disasters between today and maturity have a negligible chance
- interesting: utilizes localization to return dynamics of particular AD securities
- for a rejection of rationality, it requires establishing what negligible chance means in the model and data

In general, ϕ cannot move systematically with π_t^* .

$$\frac{\pi_t^*}{1 - \pi_t^*} = \underbrace{\frac{E_t \left[M_T/M_t \mid R_T^m = \text{low}\right]}{E_t \left[M_T/M_t \mid R_T^m = \text{high}\right]}}_{\text{constant SDF slope } \phi} \frac{\pi_t}{1 - \pi_t}$$

Inference with constant ϕ

· bad times \implies observed $\uparrow \pi_t^* \implies$ infer required increase in π_t , given ϕ

Systematic variation in ϕ

- bad times \implies observed $\uparrow \pi_t^* \implies$ associated increase in $\phi \implies$ lower required increase in π_t
- lower average ϕ needed to dampen implied variation in π_t sufficiently to satisfy the martingale restriction

Nonseparable preferences (Epstein–Zin) with interesting state dynamics Explicit time-variation in risk aversion

Models with agent heterogeneity (e.g., in risk aversion)

- · implied risk aversion a function of wealth distribution
- · how to interpret results in models with heterogeneous beliefs?

Models with financial constraints

• implied risk aversion changes with tightness of constraints

Habit formation models (Campbell and Cochrane (1999))

- implied risk aversion function of habit level
- \cdot authors show that in CC (1999), the bound still holds approximately
 - $\cdot\,$ but SDF slopes implied by the model are trivial for non-extreme states

$$\frac{\pi_t^*}{1 - \pi_t^*} = \underbrace{\frac{E_t \left[M_T / M_t \mid R_T^m = \text{low} \right]}{E_t \left[M_T / M_t \mid R_T^m = \text{high} \right]}}_{\text{time-varying SDE slope } \phi_t} \frac{\pi_t}{1 - \pi_t}$$

Question: How much variation in ϕ_t is needed such that implied π_t satisfy the martingale restriction?

• a test of a larger class of (more interesting) models

Existing models may well still fail

- many asset pricing models designed to match unconditional risk premia and the cross-section
- appropriate time-variation harder to achieve

OTHER CONSIDERATIONS

Implementation relies purely on option prices

- studies option price dynamics (with all their specifics) relative to the index
- what about correct beliefs in the pricing of the index itself?
- Authors relate identified excess belief movement to macro statistics
 - why not survey data directly?

Why bounds at all?

• Given the information obtained from localization, one could directly attempt to estimate the SDF nonparametrically (Christensen (2017))

Approximation details (discretization, etc.)

• authors take these seriously, but details are subtle

Localization

• we do not need to test adjacent states only, additional information in alternative state combinations

Given stated assumptions, the paper is cleanly executed, with a lot of detail.

- robustness checks, careful econometrics
- models could be studied in more detail

Interesting results

- much higher high frequency (daily) variation than variation at 'macro' frequencies
 - Market microstructure effects in daily data? Different models needed?
 - · Caution regarding bid-ask bounce? (not covered here)
- localization of return dynamics can discard some models
 - high AD price variation for high-return states speaks against disaster-type models that focus risk adjustments to adverse states

CONCLUSION II.

A test of rational expectations?

- $\cdot\,$ restricts attention to a class of models that is too narrow
- reverting the problem and studying required variation in implied risk aversion may be more fruitful

Even if we impose stringent restrictions on SDF and reject correct beliefs, the puzzle still stands.

- why do beliefs fluctuate so much?
- need to devise interesting models of belief dynamics to make progress ...
- ... just like interesting models of risk adjustments that go beyond imposed restrictions ...
- ... and in some cases, the two cannot be distinguished using asset price data anyway
 - e.g. Epstein–Zin with unitary IES vs. Hansen–Sargent robust preference vs. ex-post Bayesian worst-case belief

Ross (2015) 'Recovery Theorem'

- \cdot uses only cross-sectional price information, no time series
- \cdot imposes separable marginal utility (transition independence (TI))

$$m_{t+1} = \beta \frac{u'\left(\mathsf{S}_{t+1}\right)}{u'\left(\mathsf{S}_{t}\right)}$$

and stationarity of the state dynamics

- under TI, it recovers subjective belief \widehat{P} (possibly different from P)
- if TI does not hold, it recovers a long-run risk neutral measure (different from \widehat{P} or P)

This paper

- imposes a restriction based on time-series information in *P* (time-variation in prices)
- asks whether imposing $P = \hat{P}$ can be consistent with SDFs that satisfy CTI
- allows for nonstationarity of the state and hence for some martingale components in SDF
- · both papers assume time-invariant local curvature over terminal states